

# Concentration analysis of multivariate elliptic diffusions

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## Setup and aims of the paper

We consider a (weak) solution  $X$  of the SDE

$$dX_t = b(X_t) dt + \sigma(X_t) dW_t,$$

$b \in \text{Lip}_{\text{loc}}(\mathbb{R}^d; \mathbb{R}^d)$ ,  $\sigma \in \text{Lip}(\mathbb{R}^d; \mathbb{R}^{d \times d})$  and bounded, uniformly elliptic diffusion matrix  $a := \sigma\sigma^\top$ , given the drift condition

$$\langle b(x), x/\|x\| \rangle \leq -r\|x\|^{-q}, \quad \|x\| > A, \quad q \in [-1, 1]. \quad (\mathcal{D}(q))$$

The smaller the parameter  $q$ , the faster the (sub)exponential speed of ergodicity towards the invariant distribution  $\pi$  of  $X$ .

We quantify the concentration around the ergodic mean  $\pi(f) := \int f(x) \pi(dx)$  of both the continuous-time MC estimator and its discrete version

$$\widehat{\pi}_T(f) := \frac{1}{T} \int_0^T f(X_t) dt, \quad \widehat{\pi}_{n,\Delta}(f) := \frac{1}{n} \sum_{k=1}^n f(X_{k\Delta}),$$

for polynomially growing functions  $f$ .

## Strategy: Martingale approximation

For the generator  $L = b^\top \nabla + \frac{1}{2} \sum_{i,j} a_{i,j} \partial_{x_i} \partial_{x_j}$  the Poisson equation

$$Lg = f$$

has a nice solution  $L^{-1}[f]$  if  $f$  is  $\pi$ -centered and grows polynomially.

The Itô–Krylov formula then gives

$$\begin{aligned} & \int_0^t f(X_s) ds \\ &= \underbrace{\int_0^t (-\nabla L^{-1}[f](X_s))^\top \sigma(X_s) dW_s}_{(\text{loc. martingale})} + \underbrace{L^{-1}[f](X_t) - L^{-1}[f](X_0)}_{\text{remainder}} \end{aligned}$$

use polynomial growth bounds on  $L^{-1}[f], \nabla L^{-1}[f]$  from [4]

precise understanding of the tails of  $\pi$  given  $(\mathcal{D}(q))$  then allows to control all  $L^p$ -norms of the centered statistic  $\widehat{\pi}_T(f) - \pi(f)$

## Main results

**Theorem 1.** Assume  $(\mathcal{D}(q))$ ,  $\|b(x)\| \lesssim 1 + \|x\|^\kappa$  and  $|f(x)| \leq \mathfrak{L}(1 + \|x\|^\eta)$ . Let

$$\rho(\eta, \kappa, q) := \begin{cases} 1/(1-q_+), & \eta = 0 \\ \frac{1}{2} + \frac{\eta+\kappa+1+q}{1-q_+}, & \eta > 0. \end{cases}$$

Then, there exists a constant  $c > 0$  s.t. for any  $x \geq 2/\sqrt{T}$ ,

$$\mathbb{P}^\pi(|\widehat{\pi}_T(f) - \pi(f)| > x) \leq \exp\left(-c\left(\frac{x\sqrt{T}}{\mathfrak{L}}\right)^{1/\rho(\eta, \kappa, q)}\right).$$

This translates into the following PAC-bounds

Poincaré, $\eta = 0$	log-Sobolev, $\eta \leq 2$	subexponential, $\eta > 0$
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$\frac{\log(1/\delta)}{\varepsilon}$	$\frac{\log(1/\delta)}{\varepsilon}$	$\frac{\log(1/\delta)^{2\rho(\eta, \kappa, q)}}{\varepsilon^2}$
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sample complexity s.t.  $(\varepsilon, \delta)$ -PAC-bound  $\mathbb{P}^\pi(|\widehat{\pi}_T(f) - \pi(f)| \leq \varepsilon) \geq 1 - \delta$  holds

By relating the discrete estimator with sampling frequency  $\Delta$  to the continuous estimator, we obtain concentration bounds for  $\widehat{\pi}_{n,\Delta}(f)$ :

**Theorem 2.** Assume  $(\mathcal{D}(q))$ ,  $\|b(x)\| \lesssim 1 + \|x\|^\kappa$  and  $\|D^k f(x)\| \lesssim 1 + \|x\|^{\eta k}$ ,  $k = 0, 1, 2$ . Define  $\alpha := (\kappa + \eta_1) \vee \eta_2$ , and let  $\tilde{\gamma} > 1 + q$ ,  $r > 1$ , s.t.  $\tilde{\gamma} - (1 + q) > r(\alpha \vee (1 + q))/(r - 1)$ . Then, for  $p \geq 2$ ,

$$\begin{aligned} & \|\widehat{\pi}_{n,\Delta}(f) - \pi(f)\|_{L^p(\mathbb{P}^\pi)} \\ & \leq \mathfrak{D} \left( \underbrace{\Delta + \sqrt{\frac{\Delta}{n}} p^{\frac{\max\{(\tilde{\gamma}+2\alpha+1-q_+)/2, \eta_1+1-q_+\}}{1-q_+}}}_{\text{sampling error}} + \frac{1}{\sqrt{n}\Delta} p^{\frac{1+\eta+\kappa+1+q}{1-q_+}} \right) \\ & =: \Phi(n, \Delta, p), \end{aligned}$$

and

$$\mathbb{P}^\pi(|\widehat{\pi}_{n,\Delta}(f) - \pi(f)| > \epsilon \Phi(n, \Delta, x)) \leq e^{-x}, \quad x \geq 2.$$

## Sample complexities for heavy-tailed Langevin MCMC

Under mild assumptions on the potential  $U$ , the Langevin diffusion

$$dX_t = -\nabla U(X_t) dt + \sqrt{2} dW_t,$$

has invariant density  $\pi(x) \propto \exp(-U(x))$

☞ approximate sampling from  $\pi$  by numerical approximation of  $X$ , e.g., Euler scheme with  $(\xi_n) \sim_{\text{iid}} \mathcal{N}(0, \mathbb{I}_d)$ :

$$\vartheta_{n+1}^{(\Delta)} = \vartheta_n^{(\Delta)} - \Delta \nabla U(\vartheta_n^{(\Delta)}) + \sqrt{2\Delta} \xi_{n+1}, \quad \vartheta_0^{(\Delta)} \sim X_0.$$

☞ abundant literature on sampling precision in TV or Wasserstein distance for  $U$  strongly convex or modifications thereof [1, 2, 3]  
 ↗  $\pi(x) dx$  is sub-Gaussian

Assume instead that for  $q \in (0, 1)$

$$\langle \nabla U(x), x/\|x\| \rangle \geq r\|x\|^{-q}, \quad \|x\| > A. \quad (\mathcal{U}(q))$$

↗ prototypical example:  $\pi(x) \propto \exp(-\beta\|x\|^{1-q})$  is heavy-tailed and  $U$  is non-convex

## Proposition 3.

	step length $\Delta$	sample size $n$	burn-in $m$
$\varepsilon$ -prec. sampling	$\frac{\varepsilon^2}{d(\log(\mathfrak{C}/\varepsilon))^{(1-q)/(1+q)}}$	$\frac{d(\log(\mathfrak{C}/\varepsilon))^{2(1-q)/(1+q)}}{\varepsilon^2}$	–
$(\varepsilon, \delta)$ -PAC bound	$\frac{(\delta\varepsilon)^2}{d(\log(1/\delta))^{2(\eta_0+(q+3)/2)/(1-q)}}$	$\frac{d\mathfrak{D}^2(\log(1/\delta))^{(4(\eta_0+(q+3)/2))/(1-q)}}{\delta^2\varepsilon^4}$	$\frac{d(\log(1/\delta))^{2(\eta_0+q+2)/(1-q)}}{(\delta\varepsilon)^2}$

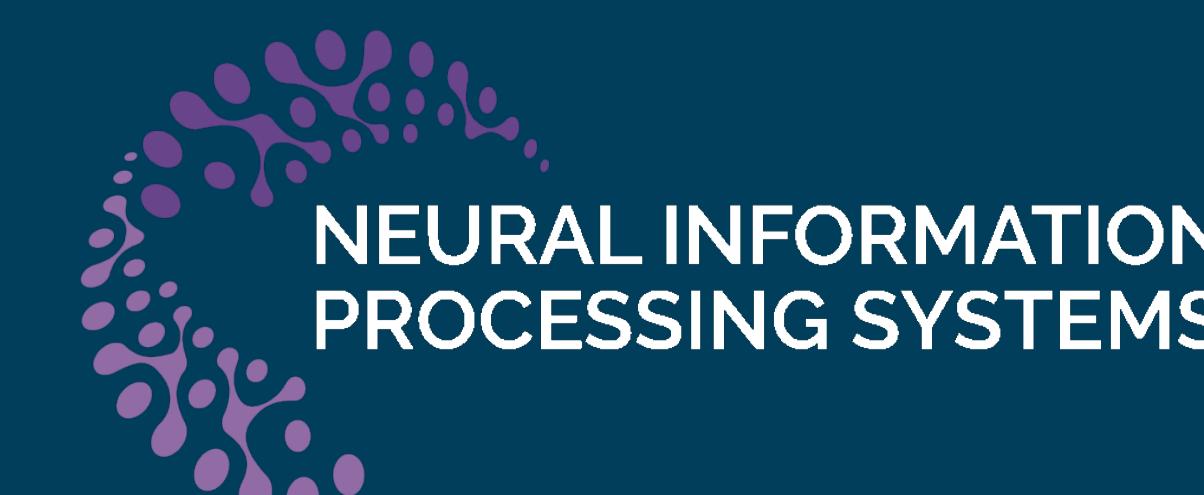
Order of sufficient sampling frequency  $\Delta$ , sample size  $n$  and burn-in  $m$  for  $(\varepsilon, \delta)$ -PAC bounds and sampling within  $\varepsilon$ -TV margin

## References

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